

4.1 By definition,  $F(y) = P(Y \leq y)$  for  $y = 1, 2, 3, \dots$

Then

$$\begin{aligned} P(Y = y) &= P(Y \leq y) - P(Y \leq y-1) \\ &= F(y) - F(y-1) \quad y = 2, 3, \dots \end{aligned}$$

$$\begin{aligned} \text{Also, } P(Y = 1) &= P(Y \leq 1) \\ &= F(1). \end{aligned}$$

4.2 a.  $F(i) = P(Y \leq i) = \sum_{k=1}^i q^{k-1} p \quad i = 0, 1, 2, \dots$

$$\begin{aligned} &= p \sum_{k=0}^{i-1} q^k = p \left( \frac{1-q^i}{1-q} \right) = \frac{p(1-q^i)}{p} \\ &= 1 - q^i. \end{aligned}$$

Because  $Y$  is a discrete random variable, the only changes in  $F(y)$  are at the positive integers. The result follows.

b. 1.  $F(y) = 0$  for  $y < 0$ . Hence,

$$\lim_{y \rightarrow -\infty} F(y) = 0.$$

2.  $F(y) = 1 - q^i \quad i \leq y < i+1$  where  $i = 0, 1, 2, \dots$

$$\begin{aligned} \text{Then } \lim_{y \rightarrow \infty} F(y) &= 1 - \lim_{i \rightarrow \infty} q^i \quad \text{for } i \text{ an integer and } 0 < q < 1 \\ &= 1 \end{aligned}$$

3. Suppose  $i \leq y_1 < y_2 < i+1$  for  $i = 0, 1, 2, \dots$

$$\text{Then } F(y_1) = 1 - q^i = F(y_2).$$

On the other hand, suppose

$$i-1 \leq y_1 < i \leq y_2 < i+1. \text{ Then}$$

$$F(y_1) = 1 - q^{i-1} < 1 - q^i = F(y_2).$$

4.8 a.  $f(y)$

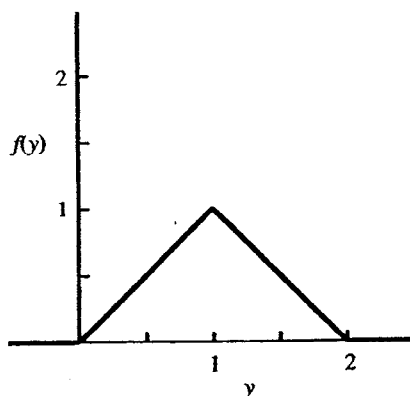


Figure 4.3

b. For  $y < 0$ ,  $F(y) = 0$ .

For  $y > 2$ ,  $F(y) = 1$ .

For  $0 \leq y \leq 1$ ,

$$F(y) = \int_0^y t \, dt = \frac{y^2}{2}$$

For  $1 \leq y \leq 2$ ,

$$F(y) = \int_0^1 t \, dt + \int_1^y (2-t) \, dt = \frac{1}{2} + \left[ 2t - \frac{t^2}{2} \right]_1^y = 2y - \frac{y^2}{2} - 1$$

c.  $P(.8 \leq Y \leq 1.2) = F(1.2) - F(.8) = (2.4 - .72 - 1) - .32 = .36$

d.  $P(Y > 1.5 | Y > 1) = \frac{P(Y > 1.5)}{P(Y > 1)} = \frac{1 - (3 - 1.125 - 1)}{\frac{1}{2}} = \frac{1.125}{.5} = .25$

VI

PROBLEM SET

4.11 a.  $F(\infty) = \int_{-\infty}^{\infty} f(y) dy = \int_0^1 (cy^2 + y) dy = c \left[ \frac{y^3}{3} \right]_0^1 + \left[ \frac{y^2}{2} \right]_0^1 = \frac{c}{3} + \frac{1}{2} = 1$

Hence  $\frac{c}{3} = \frac{1}{2}$  and  $c = \frac{3}{2}$ .

b.  $F(y) = \int_{-\infty}^y f(t) dt = \int_0^y \left( \frac{3}{2} t^2 + t \right) dt = \left[ \frac{t^3}{2} \right]_0^y + \left[ \frac{t^2}{2} \right]_0^y = \frac{y^3}{2} + \frac{y^2}{2}$  for  $0 \leq y \leq 1$

and  $F(y) = 0$  for  $y < 0$ ,  $F(y) = 1$  for  $y > 1$ .

c. The graphs of  $F(y)$  and  $f(y)$  are shown in Figures 4.6 and 4.7.

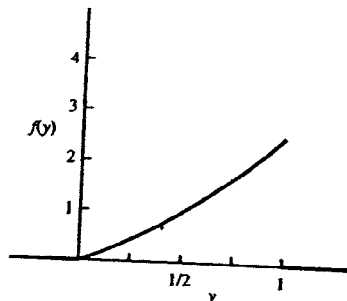


Figure 4.6

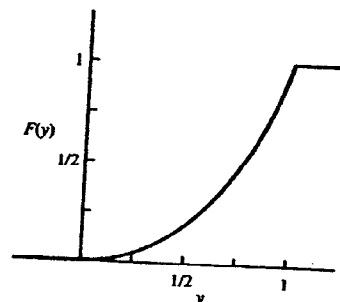


Figure 4.7

d.  $F(-1) = 0$  since  $y < 0$ ;  $F(0) = 0$ ;  $F(1) = \frac{1}{2} + \frac{1}{2} = 1$

e.  $P(0 \leq Y \leq .5) = F(.5) - F(0) = \frac{(.5)^3}{2} + \frac{(.5)^2}{2} - 0 = \frac{1}{16} + \frac{1}{8} = \frac{3}{16}$

f.  $P\left(Y > \frac{1}{2} \mid Y > \frac{1}{4}\right) = \frac{P(Y > \frac{1}{2})}{P(Y > \frac{1}{4})} = \frac{1 - \left(\frac{1}{16}\right)}{1 - \left(\frac{1}{128} + \frac{1}{32}\right)} = \frac{\frac{15}{16}}{\frac{127}{128}} = \frac{104}{123}$

4.15 Refer to Exercise 4.11.

$$E(Y) = \int_0^1 \left( \frac{3}{2} y^3 + y^2 \right) dy = \left[ \frac{3}{8} y^4 + \frac{y^3}{3} \right]_0^1 = \frac{3}{8} + \frac{1}{3} = \frac{9+8}{24} = \frac{17}{24} = .708$$

$$E(Y^2) = \int_0^1 \left( \frac{3}{2} y^4 + y^3 \right) dy = \left[ \frac{3}{10} y^5 + \frac{1}{4} y^4 \right]_0^1 = \frac{3}{10} + \frac{1}{4} = .55$$

so that  $V(Y) = E(Y^2) - (EY)^2 = .55 - (.708)^2 = .0487$ .

4.24 a.  $E(Y) = \int_0^1 y(2y) dy = \frac{2y^2}{3} \Big|_0^1 = \frac{2}{3}$

$$V(Y) = E(Y^2) - [E(Y)]^2 = \left[ \int_0^1 y^2(2y) dy \right] - \left( \frac{2}{3} \right)^2 = \left( \frac{1}{2} y^4 \right) \Big|_0^1 - \left( \frac{4}{9} \right) = \frac{1}{18}$$

b.  $E(X) = E(200Y - 60) = 200E(Y) - 60 = 200 \left( \frac{2}{3} \right) - 60 = \frac{220}{3}$   
 $V(X) = V(200Y - 60) = V(200Y) = (200)^2 V(Y) = 40,000 \left( \frac{1}{18} \right) = \frac{20,000}{9}$

c. Recall Tchebysheff's theorem from exercise 1.24.

$P(\mu_x \pm k\sigma_x) \geq 1 - \left( \frac{1}{k^2} \right)$  where  $1 - \left( \frac{1}{k^2} \right) = \frac{3}{4}$ . Solving,  $k = 2$ .

The desired interval is  $\left( \left[ \frac{220}{3} \right] \pm 2\sqrt{\frac{20,000}{9}} \right) = (-20.948, 167.614)$

3

4.29 Since the parachutist is landing at a random point in the interval  $(A, B)$ , the point of landing is a continuous random variable  $Y$ , with a uniform distribution over  $(A, B)$ . Hence

$$f(y) = \frac{1}{B-A} \quad A \leq y \leq B$$

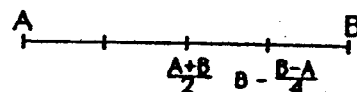


Figure 4.10

- a. Refer to Figure 4.10. If he lands closer to  $A$  than to  $B$ , he has landed in the interval  $(A, \frac{A+B}{2})$ . The probability is

$$\int_A^{(A+B)/2} \frac{1}{B-A} dy = \frac{(A+B)/2 - A}{B-A} = \frac{1}{2}$$

- b. The point at which the distance to  $A$  is exactly 3 times the distance to  $B$  is the point  $B - (\frac{1}{4})(B - A) = \frac{3B+A}{4}$ . Then

$$P(\text{distance to } A \text{ is more than 3 times distance to } B) = P\left(\frac{3B+A}{4} \leq Y \leq B\right) \\ = \frac{B - (\frac{3B+A}{4})}{B-A} = \frac{(B-A)/4}{B-A} = \frac{1}{4}$$

4.31 Recall Theorem 4.6.

$$\begin{aligned} V(Y) &= E(Y^2) - (E(Y))^2 = \left[ \int_{\theta_1}^{\theta_2} y^2 \left( \frac{1}{\theta_2 - \theta_1} dy \right) \right] - \left( \frac{\theta_2 + \theta_1}{2} \right)^2 \\ &= \left[ \frac{1}{3(\theta_2 - \theta_1)} y^3 \right]_{\theta_1}^{\theta_2} - \left( \frac{1}{4} \right) (\theta_2 - \theta_1)^2 \\ &= \left[ \frac{1}{3(\theta_2 - \theta_1)} \right] (\theta_2^3 - \theta_1^3) - \left( \frac{1}{4} \right) (\theta_2^2 + 2\theta_1\theta_2 + \theta_1^2) \\ &= \left( \frac{4}{12} \right) (\theta_2^3 - \theta_1^3) - \left( \frac{3}{12} \right) (\theta_2^2 + 2\theta_1\theta_2 + \theta_1^2) \\ &= \frac{\theta_2^3 - 2\theta_1\theta_2 + \theta_1^3}{12} = \frac{(\theta_2 - \theta_1)^2}{12} \end{aligned}$$

4.39 Let  $Y$  = cycle time. Then

$$f(y) = \frac{1}{70-50} = \frac{1}{20} \text{ for } 50 \leq y \leq 70$$

and

$$F(y) = \int_{50}^y \frac{1}{20} dt = \frac{y-50}{20} \text{ for } 50 \leq y \leq 70;$$

$$F(y) = 0 \text{ for } y < 50;$$

$$F(y) = 1 \text{ for } y > 70.$$

Thus

$$P(Y > 65 | Y > 55) = \frac{P(Y > 65)}{P(Y > 55)} = \frac{1 - \left( \frac{65-50}{20} \right)}{1 - \left( \frac{55-50}{20} \right)} = \frac{20-15}{20-5} = \frac{1}{3}$$

4.46 The next few exercises are designed to provide practice for the student in evaluating areas under the normal curve. The following notes may be of some assistance.

(4)

- (1) Table 4 tabulates the area under the standard normal curve to the right of a specified value  $z_0$ . See Figure 1. Denote the area obtained by indexing  $z = z_0$  in Table 4 by  $A(z_0)$  and the desired area by  $A$ .
- (2) Because of the symmetry of the normal distribution, and since the total area under the curve is 1, the total area lying on one side of 0 will be .5. Thus in order to calculate the area between 0 and  $z_0$  (when  $z_0 > 0$ ) we index  $z_0$ , which gives us  $A(z_0)$ . We then subtract  $A(z_0)$  from .5. That is,  $A = .5 - A(z_0)$ .

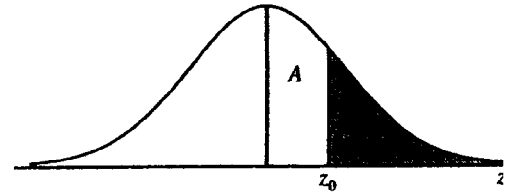


Figure 1

- (3) Notice that  $Z$  is actually a random variable that may take on an infinite number of values, both positive and negative. However, since the standardized normal curve is symmetric about 0, a left-hand area (i.e., an area corresponding to a negative value of  $z$ ) may be evaluated by indexing the corresponding positive value in Table 4.

- (a) The area between  $z = 0$  and  $z = 1.2$  is  $A_1 = .5 - A(1.2) = .5 - .1151 = .3849$ . See Figure 2.

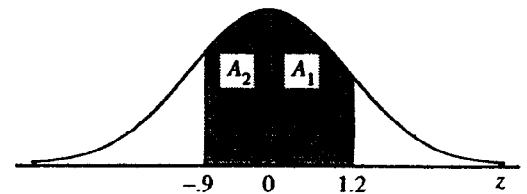


Figure 2

- (b) The area between  $z = 0$  and  $z = -.9$  is  $A_2 = .5 - A(-.9) = .5 - A(.9) = .5 - .1841 = .3159$ .

- (c) The desired area is  $A_1$ , as shown in Figure 3. Note that  $A(.3) = .3821$  and  $A(1.56) = .0594$ .  $A_1 = A(.3) - A(1.56) = .3821 - .0594 = .3227$ .

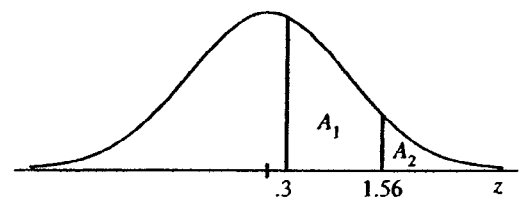


Figure 3

- (d) The desired area is  $A_1 + A_2 = .5 - A(-.2) + .5 - A(.2) = 1 - 2(.4207) = .1586$ . See Figure 4.
- (e) The desired area is  $A(-.2) - A(-1.56) = .4207 - .0594 = .3613$ .

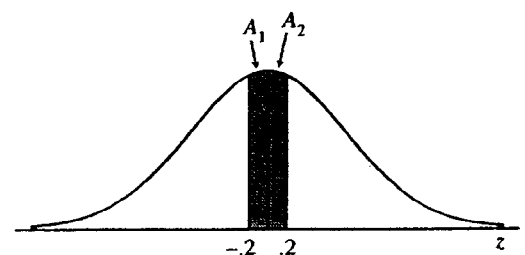


Figure 4

4.41 Let  $Y$  = time the defective board is detected.

$$a. P(0 < Y < 1) = \int_0^1 \left(\frac{1}{8}\right) dx = \frac{1}{8}$$

$$b. P(7 < y < 8) = \int_7^8 \left(\frac{1}{8}\right) dx = \frac{1}{8}$$

$$c. P(4 < Y < 5 | Y > 4) = \frac{\int_4^5 \left(\frac{1}{8}\right) dx}{\int_4^{\infty} \left(\frac{1}{8}\right) dx} = \frac{\left(\frac{1}{8}\right)}{\left(\frac{1}{8}\right)} = \frac{1}{4}$$

4.42 Let  $Y$  = amount of measurement error.  $Y$  is  $U(-.05, .05)$ .

$$a. P(-.01 < Y < .01) = \int_{-.01}^{.01} \left(\frac{1}{.1}\right) dx = .2$$

$$b. E(Y) = \frac{-0.05 + 0.05}{2} = 0$$

$$V(Y) = \frac{(-0.05 + 0.05)^2}{12} = \frac{.01}{12} = 0.00083$$

4.50 This normal distribution has  $\mu = 400$  and  $\sigma = 20$ . Probabilities associated with any normal random variable  $Y$  can be obtained by converting the necessary values of  $y$  to their corresponding  $z$  values. This conversion is made by using the formula  $z = \frac{y - \mu}{\sigma}$ . Note that  $z$  is the distance from the mean,  $y - \mu$ , measured in units of  $\sigma$ . In this case, the desired probability is  $P(Y > y_1) = 450$ . The  $z$  value corresponding to  $y_1 = 450$  is  $z_1 = \frac{450 - 400}{20} = 2.5$ . Then

$$P(Y > 450) = P(Z > 2.5) = A(2.5) = .0062$$

4.54 The fraction of students with grade point averages greater than 3.0 is given by  $A_1 = P(Y > 3.0)$  (shown in Figure 4.19). Then the  $z$  value corresponding to the point  $y = 3.0$  is

$$z = \frac{y - \mu}{\sigma} = \frac{3.0 - 2.4}{.8} = .75$$

Hence

$$A_1 = P(Y > 3.0) = P(Z > .75) = A(.75) = .2266$$

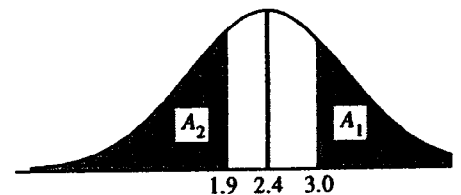


Figure 4.19

4.55 The probability of interest is  $P(Y < 1.9)$  with corresponding  $z$  value

$$z = \frac{y - \mu}{\sigma} = \frac{1.9 - 2.4}{.8} = -.625$$

(Recall that a negative value of  $z$  implies a value to the left of the mean.) Then

$$A_2 = P(Y < 1.9) = P(Z < -.625) = P(Z > .625) = A(.625) = .2660$$

(after interpolating).

4.56 Let  $X$  be the number of students with a grade point average in excess of 3.0 when 3 students are randomly selected. Then  $X$  has a binomial distribution with  $n = 3$  and  $p = P(\text{student's GPA exceeds } 3.0) = .2266$ , from Exercise 4.54. The probability of interest is

$$P(X = 3) = \binom{3}{3} p^3 q^0 = (.2266)^3 = .0116$$